## VARIOUS POINTS OF VIEW ON SOURCES

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IN ANALYTIC INFORMATION THEORY

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# APPLICATIONS to PROBABILISTIC ANALYSES of DICTIONARY STRUCTURES 

Brigitte Vallée<br>GREYC (CNRS and University of Caen)

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Two main objects: sources and data structures

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> Plan of the talk.

## Two main objects: sources and data structures

- Describe a modelling of natural sources
- Deduce consequences for the analysis of related data structures
Plan of the talk.
- A general model of sources
- The two digital structures : trie and dst.
- Probabilistic analysis of the structures, and its two steps
- Probabilistic analysis : the combinatorial step
- Probabilistic analysis : the analytic step - Need of more regular sources.
- Analysis of data structures : the result.

Sources (I)

In information theory, a source:=
a probabilistic mechanism which produces symbols from alphabet $\Sigma$, one at each time unit.
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Compromise: Only the positive part of the history is "shown"....
The negative part of the history

- is produced
- may have an influence on the positive part
- but remains "hidden"

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An origin for time : $t=0$
a sequence of random variables ( $X_{0}, X_{1}, \ldots, X_{n}, X_{n+1} \ldots$ )


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A general source may have many, strong correlations between its symbols. For $w \in \Sigma^{\star}, p_{w}:=$ probability that a word begins with the prefix $w$.

The set $\left\{p_{w}, \quad w \in \Sigma^{\star}\right\}$ defines the source $\mathcal{S}$.

A main analytical object related to any source: the Dirichlet generating functions of the source

$$
\Lambda(s):=\sum_{w \in \Sigma^{\star}} p_{w}^{s}, \quad \Lambda^{[k]}(s)=\sum_{w \in \Sigma^{k}} p_{w}^{s}, \quad\left[\Lambda=\sum_{k \geq 0} \Lambda^{[k]}\right]
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Remark: $\Lambda^{[k]}(1)=1$ for any $k, \quad \Lambda(1)=\infty$.

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- they intervene in probabilistic analyses of algorithms and data structures.

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These nice expressions are due to multiplicative properties of probabilities.
And for a general source?
Does $\Lambda(s)$ admit a nice alternative expression?

## A general source and its shifted sources

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For $u \in \Sigma^{\star}$ with $p_{u} \neq 0$, the source $\mathcal{S}_{(u)}=\left.\mathcal{S}\right|_{u}$ is a shifted source

- which gathers all the words of $\mathcal{S}$ which begin with $u \in \Sigma^{*}$,
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The conditional probabilities $p_{w \mid u}=p_{(u, v)} / p_{u}$ are denoted as $q_{v \mid u}$. These are the fundamental probabilities of the source $S_{(u)}$.

The generalized transition matrix of a source $\mathcal{S}$
The weighted graph of the source

- vertices $=$ sources $\mathcal{S}_{(u)}$
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$\mathbf{P}=$ the transition matrix of the graph.
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The non zero elements at the row $w$ are located at the columns $w \cdot i$.
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For $s \in \mathbb{C}$, the matrix $\mathbf{P}_{s}$ is obtained from $\mathbf{P}$ by raising its elements to the power $s$

## The pruned graph and the pruned matrix (I)

Sometimes, the graph (and thus the matrix) can be pruned:
With an equivalence relation on the "shifted" sources

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\mathcal{S}_{(u)} \equiv \mathcal{S}_{(v)} \quad \Longleftrightarrow \quad \forall w \in \Sigma^{\star}, \quad q_{w \mid u}=q_{w \mid v}
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one only keeps the sources $S_{(u)}$ which have a different distribution
For simple sources, this provides a finite graph (a finite matrix).


The pruned graph and the pruned matrix (II)
There are pruned graphs which remain infinite.
An instance of a VLMC (Variable Length Markov Chain)
The distribution of $X_{n}$ depends on the length of the run $0^{k}$ which precedes it


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Pruned graph :

- vertices $\mathcal{S}_{(\epsilon)}, \mathcal{S}_{(1)}$ and $\mathcal{S}_{\left(0^{k}\right)}$ for $k>0$
- all the edges labeled with 1 return to the source $\mathcal{S}_{(1)}$.

Return to the Dirichlet generating function of the source, $\Lambda(s):=\sum_{w \in \Sigma^{\star}} p_{w}^{s}$

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A general source, with its (pruned) transition matrix $\mathbf{P}_{s}$,

$$
\Lambda(s)={ }^{t} \mathbf{E} \cdot\left(I-\mathbf{P}_{s}\right)^{-1}[\mathbf{1}] \quad \text { with }{ }^{t} \mathbf{E}:=(1,0,0 \ldots)
$$

(II) Two data structures: trie and dst
dst : digital search tree - trie: shorthand for tree retrieval
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dst : digital search tree - trie: shorthand for tree retrieval

Dynamical data structures which contain words.

- Useful for sorting, and searching words.
- Important to analyze their probabilistic shape when built on a sequence of words emitted by a general source


Two types of fundamental digital structures.
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Both are trees used as dictionaries, with three main operations (Search, Insert and Delete)
Play a central role in the Lempel-Ziv data compression scheme
These trees direct words to subtrees according to their first symbol
In a trie, - internal nodes do not contain data,

- the order of insertion does not intervene.

In a dst, a word is placed on the first free node.
In a trie, the word is placed when it is alone in its subtree.
$s_{1}=b b a b \cdots ; s_{2}=a b b a a \cdots s_{3}=b a b b a \cdots, s_{4}=a b a b b \cdots ; s_{5}=b a b a b \cdots ; s_{6}=a a a a b \cdots$


## Their recursive definitions

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\operatorname{trie}(\mathcal{Y})
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- If $|\mathcal{Y}|=0, \operatorname{trie}(\mathcal{Y})=\varnothing$.
- If $|\mathcal{Y}|=1, \operatorname{trie}(\mathcal{Y})=\square$
- If $|\mathcal{Y}| \geq 2$,

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\operatorname{trie}(\mathcal{Y})=\left\langle\bullet, \operatorname{trie}\left(\mathcal{Y}_{(a)}\right), \operatorname{trie}\left(\mathcal{Y}_{(b)}\right)\right\rangle
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dst( $\mathcal{Y}$ )
- If $|\mathcal{Y}|=0, \operatorname{dst}(\mathcal{Y})=\varnothing$
- If $|\mathcal{Y}| \geq 1, \underline{\mathcal{Y}}:=\mathcal{Y} \backslash\{\operatorname{First}(\mathcal{Y})\}$
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- We will use these recursive definitions to write systems of equations.


## Role of the dst in the Lempel-Ziv Algorithm.

The Lempel-Ziv algorithm is a dictionary-based scheme

- it partitions a sequence into phrases of variable size
- a new phrase is the shortest substring not seen in the past as a phrase obtained by adding a new symbol to a "Déjà Vu" phrase


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The phrases are inserted in a DST

## Parameters for digital trees.

Two types of nodes in the digital structures

- nodes containing data - or nodes containing no data

A full node is a node containing data
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The level of a node: the length of the path from the root to it.
The size is the number of full nodes.
The two main shape parameters:

- Profile $b_{n, k}:=$ the number of full nodes at level $k$ in a tree of size $n$.
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$$
\begin{aligned}
& b_{9,0}=0 \\
& b_{9,1}=0 \\
& b_{9,2}=2 \\
& b_{9,3}=1 \\
& b_{9,4}=2 \\
& b_{9,5}=4
\end{aligned}
$$



$$
D_{n}=(1 / 9)[2 \cdot 2+3 \cdot 1+4 \cdot 2+5 \cdot 4]=3.88 \quad D_{n}=(1 / 9)[1 \cdot 2+2 \cdot 3+3 \cdot 2+4 \cdot 1]=2
$$

(III) Probabilistic analysis of the data structures.

## Probabilistic study

Input $=$ a sequence $\mathcal{X}$ of words (independently) produced by the source.
Set of inputs $=$ the set $\mathcal{M}^{\star}$ of such sequences $\mathcal{X}$
Aim $=$ the probabilistic shape of Tree $(\mathcal{X})$ for $\mathcal{X} \in \mathcal{M}^{\star}$

## Two different probabilistic models: Poisson and Bernoulli

- In the Bernoulli model, the cardinality $N$ of $\mathcal{X}$ is fixed.
- In the Poisson model, the cardinality $N$ follows a Poisson law of parameter $z$

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\operatorname{Pr}[N=k]=e^{-z} \frac{z^{k}}{k!}
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Thus: begin in the Poisson model and then return to the Bernoulli model...

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The Poisson model is easier to deal with (independence properties).
Thus: begin in the Poisson model and then return to the Bernoulli model...
For a random variable $R$ defined on the set $\mathcal{M}^{\star}$ of inputs,
there is a relation between the two expectations
$P(z)$ in the Poisson model and $B_{n}$ in the Bernoulli model,

$$
P(z)=e^{-z} \sum_{n \geq 0} B_{n} \frac{z^{n}}{n!}
$$

Two steps in the analysis of the profile polynomial $b_{N}(u):=\sum_{k \geq 0} b_{N, k} u^{k}$,
Deal with the expectations of $b_{N}(u): \quad B_{n}(u)$ [Bernoulli] and $P(z, u)$ [Poisson].
(A) The first (combinatorial) step provides an exact expression for $B_{n}(u)$

| Expectation $P(z, u)$ |
| :--- |
| in the Poisson model |$\Longrightarrow$| Mellin transform |
| :--- |
| $s \mapsto Z(s, u)$ |
| of $z \mapsto P(z, u)$ |$\Longrightarrow \quad$| Binomial expression of |
| :--- |
| the expectation $B_{n}(u)$ |
| in the Bernoulli model |

$$
B_{n}(u)=\sum_{\ell=2}^{n}(-1)^{\ell}\binom{n}{\ell} \Delta(\ell, u), \quad \text { with } \quad \Delta(s, u):=\frac{1}{\Gamma(-s)} Z(-s, u)
$$

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$$

$(B)$ The second (analytic) step provides an asymptotic estimate for $B_{n}(u)$.

- It transforms the binomial expression into an integral expression.
- It transfers the knowledge about singularities of $s \mapsto \Delta(s, u)$ into asymptotic estimates of $B_{n}(u)$
- It depends on the "tameness" of $s \mapsto \Delta(s, u)$.
(III) Probabilistic analysis : the combinatorial step.


## Profile in the Poisson model

Associate with a source $\mathcal{S}$ all its shifted sources $\mathcal{S}_{(w)}$.
Profile $b_{N, k}^{(w)}:=$ the number of full nodes at level $k$ of a digital tree of size $N$ built on the source $\mathcal{S}_{(w)}$

For a trie of size $N$


$$
\begin{gathered}
b_{N, k}^{(w)}=b_{N_{0}, k-1}^{(w \cdot 0)}+b_{N_{1}, k-1}^{(w \cdot 1)} \\
b_{N}^{(w)}(u)=u b_{N_{0}}^{(w \cdot 0)}(u)+u b_{N_{1}}^{(w \cdot 1)}(u)
\end{gathered}
$$

$$
N=N_{0}+N_{1}
$$

The number $N_{j}$ of nodes in the $j$-th subtree (that begin with the symbol $j$ ) follows a Poisson law of parameter $q_{j \mid w} z$

## System of equations on Poisson expectations.

$$
\left\{\begin{array}{lll}
P^{(w)}(z, u)=z\left(1-e^{-z}\right)+u \sum_{i \in \Sigma} P^{(w \cdot i)}\left(q_{i \mid w} z, u\right) & \text { [for trie] } \\
P^{(w)}(z, u)+\frac{d}{d z} P^{(w)}(z, u)=z+u \sum_{i \in \Sigma} P^{(w \cdot i)}\left(q_{i \mid w} z, u\right) & \text { [for dst] }
\end{array}\right.
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\end{array}\right.
$$

For each type of tree, a system of functional equations that involves in both cases

- the mapping $z \mapsto q z \quad$ - the shift on words $w \mapsto w \cdot i$
- the derivation $d / d z$ occurs for dst, not for tries.
$\Longrightarrow$ Analysis is more involved for dst.


## The Mellin transform of the Poisson expectation.

Begin with the equations satisfied by the Poisson expectations,

$$
\left\{\begin{array}{lll}
P^{(w)}(z, u) & =z\left(1-e^{-z}\right)+u \sum_{i \in \Sigma} P^{(w \cdot i)}\left(q_{i \mid w} z, u\right) & \text { [for trie] } \\
P^{(w)}(z, u) & +\frac{d}{d z} P^{(w)}(z, u)=z+u \sum_{i \in \Sigma} P^{(w \cdot i)}\left(q_{i \mid w} z, u\right) & \text { [for dst] }
\end{array}\right.
$$

Consider

- their Mellin transforms $Z^{(w)}(s, u):=\int_{0}^{+\infty} P^{(w)}(x, u) x^{s-1} d x$
- then $\Delta^{(w)}(s, u):=\frac{1}{\Gamma(-s)} Z^{(w)}(-s, u)$,
- then the vector $\boldsymbol{\Delta}(s, u)$ whose components are $\Delta^{(w)}(s, u)$.

We finally obtain a linear system for $\boldsymbol{\Delta}(s, u)$
which involves the transition matrix $\mathbf{P}_{s}$ of the source

$$
\left\{\begin{array}{cll}
\boldsymbol{\Delta}_{T}(s, u) & -s \mathbf{1} & =u \mathbf{P}_{s} \boldsymbol{\Delta}_{T}(s, u) \\
\boldsymbol{\Delta}_{D}(s, u)-\boldsymbol{\Delta}_{D}(s+1, u) & =u \mathbf{P}_{s} \boldsymbol{\Delta}_{D}(s, u) & \text { [for trie] } \\
\text { with } \mathbf{1}={ }^{t}(1,1,1, \ldots) &
\end{array}\right.
$$

The vectors $\boldsymbol{\Delta}(s, u)$ satisfy,

$$
\begin{aligned}
\boldsymbol{\Delta}_{T}(s, u) & =s\left(I-u \mathbf{P}_{s}\right)^{-1} \mathbf{1} \\
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For dst, iterate: it appears an infinite product

$$
\mathbf{Q}(s, u):=\left(I-u \mathbf{P}_{s}\right)^{-1} \cdot \ldots\left(I-u \mathbf{P}_{s+k}\right)^{-1} \ldots
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\Delta_{T}(s, u) & =s^{t} \mathbf{E}\left(I-u \mathbf{P}_{s}\right)^{-1} \mathbf{1}, & \text { [for trie] } \\
\Delta_{D}(s, u) & ={ }^{t} \mathbf{E}\left(I-u \mathbf{P}_{s}\right)^{-1} \mathbf{Q}(s+1, u) \cdot \mathbf{Q}(2, u)^{-1} \mathbf{1} & \\
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\end{array}
$$

An exact expression for $\Delta(s, u) \Longrightarrow$ a binomial expression for $B_{n}(u)$
The end of the combinatorial step.
(IV) Probabilistic analysis : the analytic step.

Return to the operator $\mathbf{P}_{s}$ and its quasi-inverse $\left(I-u \mathbf{P}_{s}\right)^{-1}$.

## Return to the operator $\mathbf{P}_{s}$ and its quasi-inverse $\left(I-u \mathbf{P}_{s}\right)^{-1}$.

Remind: $\mathbf{P}_{s}$ is a matrix whose rows and columns are induced by $\Sigma^{\star}$. Its non zero coefficients at row $w$ are located at columns w.i, and are equal to $q_{i \mid w}^{s}=\left(p_{w \cdot i} / p_{w}\right)^{s}$

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Its non zero coefficients at row $w$ are located at columns $w . i$, and are equal to $q_{i \mid w}^{s}=\left(p_{w \cdot i} / p_{w}\right)^{s}$

The operator $\mathbf{P}_{s}$ operates on $L^{\infty}\left(\Sigma^{\star}\right)$ in a natural way:
$L^{\infty}\left(\Sigma^{\star}\right):=$ the Banach space of the bounded functions $X: \Sigma^{\star} \rightarrow \mathbb{C}$, endowed with the sup norm.

$$
Y=\mathbf{P}_{s}[X] \quad \Longleftrightarrow \quad Y(w)=\mathbf{P}_{s}[X](w):=\sum_{i \in \Sigma} q_{i \mid w}^{s} X(w \cdot i)
$$

$\mathbf{P}:=\mathbf{P}_{1}$ is stochastic, $\Longrightarrow$ a dominant eigenvalue equal to 1.
Need : precise information for the quasi-inverse $\left(I-u \mathbf{P}_{s}\right)^{-1}$
for $u$ close to 1 and $\Re s$ close to 1 .
Related to spectral properties of $\mathbf{P}_{s}$ on a convenient functional space....

## Which functional space ?

There are two cases (for the source)

- (i) The pruned graph becomes finite
- (ii) it remains infinite.

There are two cases (for the tree) $=$ the $T$-case and the $D$-case.
For (ii) - we have to find a space where the infinite matrix $\mathbf{P}_{s}$ well behaves.

- there is an extra difficulty in the $D$-case: the infinite product and we thus need a source with a past


## Sources with a past

When the symbol $X_{n}$ is emitted,

- it "looks at" (from its relative point of view) its neighbors, - which form its reverse past $X_{n-1}, \cdots, X_{1}, X_{0}$ in this order

- If $w$ is the previously emitted prefix, it considers its mirror $\phi(w)$.


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Properties of $g$ for "simple" sources:
Memoryless source $\Longleftrightarrow g$ constant on each $i \cdot \Sigma^{\star}, i \in \Sigma$
Markov chains of order $1 \Longleftrightarrow g$ constant on each $i j \cdot \Sigma^{\star}, i, j \in \Sigma$
Markov chains of order $k \Longleftrightarrow g$ constant on each $w \cdot \Sigma^{\star}, w \in \Sigma^{k+1}$

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Markov chains of order $k \Longleftrightarrow g$ constant on each $w \cdot \Sigma^{\star}, w \in \Sigma^{k+1}$
For "good" sources: one may assume $g$ to be continuous or even Hölder with respect to the usual "distance" $\delta$ on $\Sigma^{\star}$,
$\delta(x, y)=2^{-\gamma(x, y)} \quad$ where $\quad \gamma(x, y)$ the coincidence between $x$ and $y$

## Sources with an infinite past.

If the source is regular enough (with a Hölder $g$-function for instance), this finite reverse past can be extended to an infinite reverse past It admits a stationary measure, and we consider the stationary source.


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The (mirror of) transition matrix $\mathbf{P}_{s}$ is then extended into an operator - which acts on the space $\mathcal{H}\left(\Sigma^{\mathbb{N}}\right)$ of Hölder functions $X: \Sigma^{\mathbb{N}} \rightarrow \mathbb{C}$.

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- with now good spectral properties

$$
\begin{aligned}
& \left(I-u \mathbf{P}_{s}\right)^{-1} \text { is extended to }\left(I-u \mathbb{H}_{s}\right)^{-1} \text { which is "tame" } \\
& \text { for } \Re s \text { and } u \text { close to } 1 .
\end{aligned}
$$ and the $\Delta(s, u)$ related to the two data structures

$$
\begin{array}{lll}
\Delta_{T}(s, u) & =s^{t} \mathbf{E}\left(I-u \mathbf{P}_{s}\right)^{-1} \mathbf{1}, & \text { [for trie] } \\
\Delta_{D}(s, u) & ={ }^{t} \mathbf{E}\left(I-u \mathbf{P}_{s}\right)^{-1} \mathbf{Q}(s+1, u) \cdot \mathbf{Q}(2, u)^{-1} \mathbf{1} & \text { [for dst] }
\end{array}
$$

are also "tame", with a "tameness" of the same type.
(V) Probabilistic analysis : the result.

## Main results

Consider a stationary tame source $\mathcal{S}$, and a digital tree built on $n$ words independently drawn from the source. We consider a trie (type $T$ ) or a dst (type $D$ ), with $X \in\{T, D\}$

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The mean and the variance of the depth $D_{n}$ satisfy

$$
\begin{aligned}
\mathbb{E}\left[D_{n}\right] & =\mu \log n+\mu_{X}+R(n) \\
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- The dominant constants $\mu, \nu$ only depend on the source, not on the tree type
- The subdominant constants $\mu_{X}, \nu_{X}$ depend on the source and the tree

The inequality $\mu_{T}>\mu_{D}$ holds.

- The remainder terms $R(n)$ depend on the tameness of the source.


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When the source $S$ is not an unbiased memoryless source, one has $\nu \neq 0$ and the depth $D_{n}$ asymptotically follows a Gaussian law

$$
\frac{D_{n}-\mathbb{E}\left[D_{n}\right]}{\sqrt{\mathbb{V}\left[D_{n}\right]}} \xrightarrow{d} \mathcal{N}(0,1) \quad\left[\text { speed of convergence } O(\log n)^{-1 / 2}\right] .
$$

## Precise results



$$
\mathbb{E}\left[D_{n}\right]=\mu \log n+\mu_{X}+R(n), \quad \mathbb{V}\left[D_{n}\right]=\nu \log n+\nu_{X}+R(n)
$$

| Dominant terms | Types of tameness | Remainder terms |
| :--- | :---: | :--- |
| $\mu=-\frac{1}{\lambda^{\prime}(1)}$ | $S$-tame | $O\left(n^{-\delta}\right)$ |
| $\nu=\frac{\lambda^{\prime}(1)^{2}-\lambda^{\prime \prime}(1)}{\lambda^{\prime}(1)^{3}}$ | $H$-tame | $O\left(\exp \left[-(\log n)^{\rho}\right]\right)$ |
|  | $P$-tame | $\psi(n)+O\left(n^{-\delta}\right)$ |

$-\lambda(s)$ is the dominant eigenvalue of the source $\left(I-\mathbb{H}_{s}\right)^{-1} \leadsto 1 /(1-\lambda(s))$
$-\delta$ and $\rho$ : related to the geometry of the tameness
$-\psi(n)$ : a periodic function of $\log n$

## Conclusion

Description of the interaction between the source and the data structures,

- via the $\Delta(s, u)$ functions called the mixed Dirichlet series.
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Analyses of sorting or searching algorithms when they deal with words, with the cost "number of symbols that are used for comparing words".

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Description of the interaction between the source and the data structures,

- via the $\Delta(s, u)$ functions called the mixed Dirichlet series.
- precise comparison between the two structures (trie, dst).

Other instances of this interaction:
Analyses of sorting or searching algorithms when they deal with words, with the cost "number of symbols that are used for comparing words".

Open question:
Is it possible to return to the analysis of the Lempel-Ziv algorithm?

## What happens on the left of the vertical line $\Re s=1$ ?

It is important for the analysis to deal with a region $\mathcal{R}$ where $\left(I-\underline{\mathbf{P}}_{s}\right)^{-1}$ is tame : analytic (except for $s=1$ ) and of polynomial growth $(\Im s \rightarrow \infty)$

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Different possible regions $\mathcal{R}$ where $\left(I-\widehat{\mathbf{P}}_{s}\right)^{-1}$ is tame.

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Different possible regions $\mathcal{R}$ where $\left(I-\underline{\mathbf{P}}_{s}\right)^{-1}$ is tame.


Situation 1
Vertical strip
$1-\sigma \leq a$


Situation 2
Hyperbolic region
$1-\sigma \leq t^{-\alpha}$


Situation 3
Vertical strip with holes

Possible tameness regions for a simple source


Situation 1
Vertical strip


Situation 2
Hyperbolic region


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Vertical strip with holes

Possible tameness regions for a simple source


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For which simple sources do these different situations occur?

Possible tameness regions for a simple source


Situation 1
Vertical strip


Situation 2
Hyperbolic region


Situation 3
Vertical strip with holes

For which simple sources do these different situations occur?
For memoryless sources relative to probabilities $\left(p_{1}, p_{2}, \ldots, p_{r}\right)$

- S1 is impossible
- S3 occurs when all the ratios $\log p_{i} / \log p_{j}$ are rational
- S2 occurs if there exists a ratio $\log p_{i} / \log p_{j}$ which is "diophantine" [badly approximable by rationals]

For which Lipschitz, stationary, smooth sources do these different situations occur?

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Situation 1
Vertical strip
Geometric condition


Situation 2
Hyperbolic region
Arithmetic condition


Situation 3
Vertical strip with holes
Periodicity condition

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Vertical strip
Geometric condition


Situation 2
Hyperbolic region
Arithmetic condition


Situation 3
Vertical strip with holes
Periodicity condition

- S1: When ? Find some equivalent of the UNI Condition 'the branches are not too often of the same shape" (??)
- S3: only when the source is conjugated to a simple source.
- S2: when the following condition [DIOP] holds "there exists two cycles $\mathcal{C}_{i}$ and $\mathcal{C}_{j}$ for which the ratio $\log p\left(\mathcal{C}_{i}\right) / \log p\left(\mathcal{C}_{j}\right)$ is "diophantine"

